Study and Development of Amalgamating Rules for the Reliability Indicators of Power train System Elements

Belodedenko S.V.1*, Hanush V.I.1, Hrechanyi O.M.1, Ibragimov M.S.1
Chair “Machines and units of metallurgical production”, National metallurgical academy of Ukraine, Dnipro, 49600, Ukraine

1. ARES OF APPLICATION OF RULES FOR AMALGAMATING RELIABILITY INDICATORS

In terms of assessing reliability, there are two approaches. The combined approach is that the entire mechanical system as a whole is tested or observed. The mean time between failures \( t_\Sigma \) is used as the primary information, after which a statistical sample of \( n \) members is formed. It is processed by appropriate methods, after which the type of the reliability function \( P_{cl}(t) \) is established (Fig. 1). Such an approach is sometimes called the classical reliability methods [1]. It is dominated by mathematical-statistical methods, invariant to the type of failure. They are designed, first of all, for electronic systems with a large number of elements operating in parallel, subject to sudden failures. For the power mechanical systems, representing, predominantly, a series-connected structure, undergoing gradual effects of several degradation processes, these methods are poorly adaptable.

Fig1. Two approaches to assessing the reliability of systems.

The main problem of the combined approach is related to the limited sample of the number of facts of failure \( n \), which does not allow to correctly determine the form of the reliability function. Reliability
tests of full-scale mechanical systems are expensive; therefore it is advantageous to investigate reliability during operation. But at a considerable interval of system operation, fails only its non-critical elements. Although interest is information that contributes to insurance against significant losses.

Such information in sufficient quantities can be obtained at a lower level by testing small-sized structural elements. In relation to this, algorithms have been developed for obtaining of the lifetime distribution functions (LDF) of certain critical elements of a mechanical system. This is an individual approach to assessing reliability, which is also called structural reliability methods [1]. Based on the amalgamation, according to the relevant rules of the function family, the reliability function $P_{st}(t)$ is obtained (Fig. 1). In the individual approach, probabilistic-physical methods are used, taking into account the nature of the degradation processes (the failure physics) and the mutual influence of elements [2, 3].

In the modern theory of maintenance, the boundary condition is determined by the moment of transition of the system to the corresponding phase of the technical condition, the number of which quantity tends to increase [3]. The same trend is observed in terms of the number of damaging processes affecting the system.

Mathematical-statistical methods of the classical approach should be used in maintenance strategies with reliability centered maintenance. This applies to facilities of mass production, a significant number of homogeneous copies of which are operated in the same conditions. For unique production facilities, it is advisable to use a risk-based strategy. An individual approach together with probabilistic-physical methods is adequate to it. However, one should not oppose the two approaches one should skillfully use the advantages of both methods [1].

When optimizing the periodicity of inspections by the criterion of the minimum cost of operation, one should have a reliability function $P(t)$. In traditional optimization algorithms of Dhillon, Pham, Christer it is obtained by classical methods [4-6]. In this case, the function $P(t)$ is evaluated a posteriori according to the facts of system failures. The latter have a different scale for consequences, as a result of which the reliability function gets a “blurred” character. This leads to a reduction periods between inspections and over maintenance effect.

The development of technical diagnostic tools made it possible to monitor the most dangerous damage of the fatigue type. This led to the emergence of algorithms Reliability Based Optimization (RBO) [7-9]. In them, the reliability function is estimated based on the probability of defect detection. Their critical sizes and growth intensity are predicted by probabilistic-physical methods. Namely, both fracture mechanics models and fatigue resistance models ($S-N$ models) are used [10]. The effectiveness of RBO- algorithms is high if the type of cracks is known. In fact, the “weak link” principle is used here. In case of multifocal damage, the obtained reliability indices need anyway amalgamated.

The amalgamation procedure is actual when using structural reliability methods. This may be in the action of a complex of damaging processes $D_{1..k}$ on the structural element $E$ (Fig. 2, a), and in the action of a certain damaging process $D_1$ on the system of elements $E_{1..i}$ (Fig. 2, b). Each process $D_{1..k}$, each element of $E_{1..i}$ has some degree of criticality, respectively, $u_{1..k}$ and $u_{1..i}$ in terms of their failure. In the general case, according to individual reliability indices $P_{ik}$ or safety (risk) $\beta_{ik}$ receive general characteristics of the system $P_{st}(t)$ and $\beta_{st}(t)$, which change with the operating time $t$ (Fig. 2, c).
Fig 2. The scheme for solving the reliability problem $P$ and the safety $\beta$ of power mechanical systems that represent the complex of damaging processes $D$ (a), the system of series connected elements $E$ (b), as well as the scheme for the general case of the system with $D_k$ processes and $E_i$ - elements (c).

The concept of acceptable risk, on which the theory of industrial safety is based, provides for a staged reassignment of service lifetimes (updating). The question of the form of the $LDF$ and the inverse of the reliability function $P(t)$ is the key in this situation. If, in the classical approach, the reliability function is considered to be the initial index of the system (which is mostly of a posteriori nature), then in the structural approach for the initial indicator it is necessary to accept the $LDF$ of the elements having a priori nature.

The importance of amalgamating algorithms in maintenance is due to the fact that the planning of recovery operations takes place, at least, at the machine level. It is inappropriate to appoint inspections for only one element, since at that time the entire system will not be available. The amalgamation of indicators is carried out according to some rules, reviewing which, and developing new ones, is devoted to this paper.

2. **Complex Indicators of Technical Condition**

The problem of amalgamation of individual indicators is exacerbated at the stage of operation when diagnosing the technical state of the mechanical system. After all, in addition to the considered indicators of the elements, it is necessary to take into account the influence of the complex of degradation processes on them. In addition, thanks to the development of technical diagnostic systems, the number of technical indicators that are heterogeneous has increased. They are also subject to integration in the process of assessing the state of the object for the adoption of one of the three decisions. 1. Prolongation of its operation with standard parameters, or with their limitations. 2. Repairs and upgrades with subsequent use in order. 3. Decommissioning.

Combine heterogeneous non-comparable indicators of qualitative and quantitative analysis, with varying degrees of identification (uncertainty), may by methods of fuzzy logic. They are used, for example, in the Intelligent Structural Health Management of large industrial objects, which includes diagnostic systems and prognostic system. Such systems should be used at the level of the equipment of a production site.

In order to assess the technical condition of powertrain units, it is necessary to use complex (integrated) indicators. The probability of survival $P$ (PS) is used as a complex diagnostic indicator for facilities mass production. In the classical formulation, PS describes the relative number of failures. On the one hand, the prediction algorithms for this indicator should not be too sensitive to the growth of the number of technical system elements. The upward trend in the number of calculated and diagnosable elements may lead to an unjustifiably low predicted level of the PS of the entire system. As a result, the cost of the object increases. On the other hand, PS should respond to the operating time if it acts as a diagnostic parameter. Both conditions contradict each other. This circumstance motivates the search for better indicators.

First of all, the technical condition is estimated directly by the statistical reserve or reliability index $\gamma$ (its properties are discussed below). In addition, greater sensitivity is achieved by converting the probabilities of survival $P$ and failures $Q = 1 - P$ in their ratio. This technique is known as the odds ratio. The application of the logarithm of the odds ratio (logistic function) brings significant advantages in determining the technical condition.
For a particular element of the mechanical system, the PS ceases to characterize the relative number of failures. Therefore, it requires indicators that are filled with physical content. In this aspect, it is promising to obtain a diagnostic indicator related to the residual operating time (residual, remain resource).

3. RULES OF AMALGAMATING INDIVIDUAL RELIABILITY INDICES FOR THE RELIABILITY INDICATOR OF THE WHOLE POWERTRAIN UNIT

Powertrain units are used in drives of machines for the transfer of efforts, energy and its transformation. These, as a rule, include gearboxes, motors, engines transmissions, pumps. From the standpoint of reliability, the powertrain, most often, is modeled by a series of connecting elements, the failure of which leads to the loss of the available unit. Redundancy at the element level occurs, as a rule, somewhat peculiar - by the creation of statically uncertain structures. But due to the lack of calculation methods for them, it can be assumed that redundancy is carried out at the level of at least subsystems of units. Thus, powertrain units should be classified as simple technical systems, the prediction of which is possible on the principle of a "weak link". Such a conclusion cannot be fully perceived, because, first, it is impossible to accurately define a "weak" element. Only the roller bearing is breaking for almost ten reasons. For gears of such reasons are even bigger. Secondly, powertrain units can gradually lose efficiency (power, efficiency, productivity), reaching a certain limit, that can be considered a failure. Therefore, the system may function in a damaged condition. Such a situation is inherent in complex technical systems.

The development of technical diagnostics makes it possible to observe the effects of the complex of damaging processes. This ultimately increases reliability. Simple technical systems are presented as complex. This is accompanied by modernization of techniques for predicting the technical condition of objects. In such circumstances, an amalgamating the indicators of individual elements of the system into a common, generalized system indicator becomes an important process.

The most commonly used method of amalgamating the individual probabilities of the survival (PS) $P_i$ of individual elements of the system into its common PS $P_Σ$ is the rule of multiplication:

$$P_Σ = \prod_{i=1}^{n} P_i$$

In this case $n$- number quantity of system elements and degradation processes acting on them. The rule of multiplication of the PS corresponds to the rule of adding individual failures ratio: $\lambda_Σ = \sum \lambda_i$. Hence for the exponential law of reliability we obtain:

$$P_Σ = \exp(-\lambda_Σ t)$$

The widely-known disadvantage of this rule is an excessive drop in the value of $P_Σ$ with an increase in the number of elements of the system (line PI, Fig.3). It is established that calculating the reliability of systems based on the exponential distribution (lambda method) leads to a huge methodological error - to underestimate the estimation of the mean time to failure of the system in $\sqrt{n}$ times, where $n$ is the number of elements in the consecutive system.

**Fig3.** The change of the reliability in the relative operation time $\lambda t$ of the system $P_Σ$, which consists of 10 elements (their failure ratio: $\lambda_1 = \lambda_2 = \lambda_3 = 0.1$ month$^{-1}$, $\lambda_4 = \lambda_5 = \lambda_6 = 0.2$ month$^{-1}$, $\lambda_7 = \lambda_8 = \lambda_9 = \lambda_{10} = 0.3$ month$^{-1}$). Its obtained by the rules (1) (PI), (2) (PII), (3) (PIII), (4) (PIV, $z_n = 0.5$), (5) (PV, $[1-U] = 0.5$), (8) (PVI).
In this regard, the multiplication rule is used to estimate the lower bound of reliability in the first approximation [11]:

\[
\prod P_i \leq P_\Sigma \leq P_{\min}.
\] (1b)

The tighter boundaries based on the probability of compatible events taking into account the mutual correlation of failures were later developed [12].

Significant decrease in the reliability of the system relative to the reliability of its elements in the design stage provocative the rise of cost mechanical systems. This happens either through the growth of machines weight, or through the use of more high-quality and expensive materials and technologies. The high cost of the system contributes to the fact that the optimal period of its maintenance is reduced.

Practical recommendations for using the multiplication rule are established. 1. If the coefficient of variation of the lifetime elements of the system does not exceed 40%, then you can consider only 5 elements to find the value of \( P_\Sigma \). 2. It is also possible not to take into account PS elements whose durability is more than 5 times higher than the durability of the weak link [13].

With regard to the reliability of gearboxes, the upgraded formula (1) is proposed:

\[
P_{\Sigma \text{III}} = 0.14 + 0.86 \prod P_i.
\] (2)

For a definitive determination of the reliability of series systems, the "formula of chain" has become widespread:

\[
P_{\Sigma \text{III}} = \left[ \sum \frac{1}{P_i} - (n - 1) \right]^{-1}. \tag{3}
\]

The results obtained in both formulas are not very different from the dependence of \( P_\Sigma (\lambda t) \) (lines PII, PIII, Fig. 3).

The influence of the dependence between elements of the system is taken into account through the coefficient \( z_n \). In this case, the following dependence [14] applies:

\[
P_{\Sigma \text{IV}} = \prod P_i + z_n (\min P_i - \prod P_i). \tag{4}
\]

The coefficient \( z_n \) increases with the increase in the number of elements of the system \( n \), and decreases with the growth of the statistical reserve. Due to this, the decrease in the level of \( P_\Sigma \) occurs less intensively. The given method flexibly regulates system's reliability (line PIV, fig.3), but requires additional research coefficient \( z_n \).

The multiplication rule gives a pessimistic estimate of predicted reliability, which, most often, is not supported by experiments. A similar phenomenon is known in the methodology of risk analysis. Its main tool is the relationship between cumulative frequency of an accident and its severity of consequences (F-N curves, Farmer's curve), has a form of graduate function. Its degree is the risk aversion factor. A situation arises when a mechanical structural does not perceive risk at the expected level, demonstrating a more optimistic scenario. In this case, the dependence is called a risk aversion curve [15]. This fact gives hope that the use of risk analysis tools will solve the problem of amalgamating the individual ability indicators of individual elements of the system into its common indicator. In its form, the formula used to determine the reliability of systems with structural redundancy is well suited for delaying the fall of the value of \( P_\Sigma \):

\[
P_{\Sigma \text{V}} = \left[ 1 + \sum k_f \frac{1 - P_i}{P_i} \right] \prod P_i. \tag{5}
\]

Here the redundancy is estimated through the coefficient \( k_f < 1 \). For a simple system of \( k_f = 0 \) and the large risk of a failure. With full redundancy of \( k_f = 1 \) and the risk of failure (consequences of failure), practically no. Unfortunately, \( k_f \) values are obtained by testing the entire system. To get rid of this, you can use the coefficients of significance (criticality) of failures that exist in the risk analysis: \( u_i = 1 - k_p \). The greater the critically of failure in its consequences, the smaller is the addition to increasing the system's minimum reliability. The fraction in (5) is a dimensionless risk index \( \rho_i \) [15]. In essence,
Study and Development of Amalgamating Rules for the Reliability Indicators of Power train System Elements

this is a odds ratio. With deterministic calculations for \( P = 0.5 \), the risk becomes complete, that is, \( \rho = 1 \). In this case (5) transforms into:

\[
P_{\Sigma V} = \left[1 + \Sigma (1 - u_i) \rho_i \right] \prod_i P_i .
\]  
(5a)

With this dependency, you can effectively influence the amalgamation process. But it is not effective at extreme critical values, i.e. for \( u_i = 0 \) and for \( u_i = 1 \).

From one the modifications to the Lindley distribution [16, 17] follows a new rule of amalgamation, which in structure is similar to the structure of formulas (5). One-dimensional form of Lindley distribution has cdf:

\[
F(t) = 1 - \frac{\theta + 1 + \theta t}{\theta + 1} \exp(-\theta t).
\]  
(6)

The scale parameter \( \theta \) can be expressed through the failure rate \( \lambda : \theta = \lambda / (1 + \lambda) \). Then equation (6) can be represented as:

\[
F(t) = \theta [1 - \exp(-\lambda t)] + (1 - \theta) [1 - (1 + \lambda t) \exp(-\lambda t)].
\]  
(7)

Hence, if by the addition rule find the total failure rate for the \( \lambda_x \) system, its reliability will be found as:

\[
P_{\Sigma V} = (1 + \lambda_x t) \exp(-\lambda_x t).
\]  
(8)

Marking \( P_{EXP} = P_t \), these formulas are transformed into \( P_{\Sigma V} = P_{IND} = (1 + \lambda t) P_{EXP} \).

From Fig.3. It is seen that the classic rule of multiplication is the most conservative estimate of the system's reliability. In the opinion of the authors, such a rule should be used at the design stage for mechanical systems consisting of elements of relatively low reliability. To predict the reliability of responsible systems with high-reliability elements it is worth using qualitatively other models, which ensure the principle \( P_x \rightarrow P_{min} \). Moreover, it is advisable to do it at the stage of operation, when some reliability indicators are already known.

As a result of valid studies, it was found that the rules for amalgamating \( P_v, P_{\Sigma V} \) correspond to the principle \( P_x \rightarrow P_{kmin} \). However, they are sensitive enough to increase the number of system elements. With their increase it is difficult to achieve the desired principle. To solve this problem, one must understand what the cause of such a phenomenon is.

The rule of multiplication of the PS is fair for independent events. If there are deviations from this rule, then it is assumed that these are the consequences of the mutual influence of the elements of the system [14]. This effect is manifested during operation. When forecasting the system's reliability, there is the amalgamation effect that involves a significant reduction of the magnitude \( P_x \) relative to the \( P_{min} \) values. The amalgamation effect is formalized, as \( P_x \ll P_{kmin} \). It is the outcome of uncertainty. Failure of the system occurs under the influence of the dominant damaging process for the most vulnerable element. Combating the effect of the amalgamation is the procedure for identifying (updating) the model of the operation process. Its result is the simplification of a complex technical system to the simple, whose robustness is valued on the principle of weak link. Thus the principle \( P_x \rightarrow P_{kmin} \) is realized.

This corresponds to the rules of combining risk indicators, one of which is the resource (lifetime) safety index (RSI) [15].

4. THE ORIGIN OF THE RESOURCE (LIFETIME) SAFETY INDEX

When diagnosing a safe state of the system is evaluated by comparing the indicators of the damaging (degradation) process \( y \), acting on the system, and indicators of system resistance to process \( Y \). In the general case, performance indicators \( Y \) and \( y \) with sufficient informativity can serve as indicators. Complex indicators, unlike simple ones, are more informative, but they are more difficult to control. Simple indicators are effective for simple technical systems by the type of "weak link".

In order to determine the probability of survival operation during sudden failures, the model of "load-strength" comparison is usually used. Reliability is defined as the probability of exceeding the strength over the load. The "load-strength" model has been worked out in detail for various combinations of statistical distributions and is a classic approach. As a result this model is also used for failures of a gradual type. But for fatigue failures the conditionally selected limit value of the load can exceed the minimum strength level in many times not reflecting on the actual reliability. But
according to the existing model, there should be a failure. To overcome this contradiction several ways not proved to be effective are offered.

The comparative model “operation (operating time)-resource (lifetime)”, which distinguishes the resource-based approach, is a comprehensive solution to the problem of gradual failures. It works fine at the operation stage, when the diagnosis of the residual resource is carried out by controlling the natural parameter, which is the operating time \( t \). The ratio of maximum operation lifetime \( T \) to the minimum resource \( T_P \) (which is determined in the statistical aspect by its own distribution functions) forms a guaranteed safe factor to operation life \( n_{TP} \). The logarithm of its current value in the form of the safety index decreases linearly with the operating time. When the RSI reaches zero value it indicates that the object is used with the unacceptable risk.

The \( n_{TP} \) value is inverse to the probabilistic accumulated damage in the resource interpretation \( a_p \).

Then for an individual safety index RSI is fair:

\[
\beta_{pk} = \log \frac{T_p}{t} \log n_{TP} = \log a_p^{-1} = \beta_{pk0} - \log t .
\]  

(9)

The similar use of the risk indicator for sudden type failures is not possible according to the classical model. In general, the model “load - strength” is not adapted for the current control of the technical condition. Therefore, the purpose of the available researches is to use the resource model for sudden failures and develop the algorithm for the safety indexes determination. Harmonization of the technical condition estimation methods as the result gives an opportunity to increase the service reliability.

4.1. Approaches to the Interpretation of the Statistic Margin.

The probability of failure is usually interpreted as the overlap area under the left branch of the pdf \( f(y) \) plot and the right branch of the pdf \( f(Y) \) plot. Therefore, the failure probability \( Q \) is determined by the statistical margin \( \gamma \) and the Laplace’s function \( \Phi \):

\[
Q = \Phi(\gamma). \quad (10)
\]

For normally distributed independent indices \( y \) and \( Y \) (for \( Y>y \)) we have:

\[
\gamma = \frac{\bar{Y} - \bar{y}}{\sqrt{S_Y^2 + S_y^2}},
\]  

(11)

where \( \bar{Y} \) and \( \bar{y} \), and \( S_Y \) and \( S_y \) are respectively the median values and the standard deviations (SD) of \( y \) and \( Y \) indices.

The statistical margin in this formulation, referred in the literature as the Cornell's safety index [18], can be interpreted as the minimum distance from the center of \( O \) to the line corresponding to the equation \( \varepsilon = \bar{Y}-\bar{y} \), separating the safe state from the failure one (fig. 4).

**Fig4. Prior to the statistic margin (right), and the correction of the value \( a \) to the next (left).**

The position of the centre \( O \) is determined by the dimensionless indices of the technical state \( \gamma_y = \bar{y}/S_y \) and \( \gamma_Y = \bar{Y}/S_Y \), which are opposite to the variation values coefficients \( \gamma \) i \( Y \). The denominator of the formula (11) is the SD of the value \( \varepsilon \). The method of statistical reserve (in the technical literature it is called first-order-second-moment method [19, 20]) is developed for multidimensional reliability situation when the object is subjected to multi degradation processes, each of them has the safety index \( \gamma_i \) obtained by (11). Then the formula (10) determining the probability of the boundary state *is converted into* \( Q = \Sigma \Phi(\gamma) \) [21]. This proves the proper use of the rule for summing up the system risks.
The described approach is suitable for sudden failures; for gradual failures it is rather conditional, and for fatigue failure estimation it is ineffective. To calculate of the reliability function of powertrain systems in the case of sudden failures, a method that takes into account the dynamic interaction between the load (parameter $y$) and strength (parameter $Y$) looks promising. In this case, instead of their fixed values, the time functions are used: $\varepsilon(t) = Y(t) - y(t)$. Due to this, you can get rid of the excessive conservative assessment of reliability inherent in consistently combined structures. But such an analysis looks complicated, despite the fact that it contains only a few primitive data [22].

Providing, that the safety index and statistical margin both are the resource interpretations, it is possible to use formula (10) for gradual failures. In this case the probability of failure is represented by comparing the operation life distribution functions. Here the probability of the failure is represented by comparing the operating time $t_P$ and resource $T_P$ distribution functions and in the general case by comparing the parameters $y_P$ and $Y_P$ distribution (Fig. 5). For its determination we use the area of positive quantiles $u_Q$ for the graph $t_P$ and $y_P$ and the area of negative quantiles - $u_Q$ for the graphs $T_P$ and $Y_P$. The probability of a failure in the form of its quantile will correspond to the cross point of these graphs (Fig. 6).

![Fig 5](image1.png)

**Fig 5.** Scheme for the formation of the RSI and the function of gradual failures in the lifetime functions of distribution of $T_P$ and operating time $t_P$.

![Fig 6](image2.png)

**Fig 6.** To determine the statistical margin when comparing the distribution functions of the diagnostic parameter

At the same place the safety index $\beta_{2R}$ equals zero. As the slope of the $t_P$ and $T_P$ graphs is determined by the SD operation life and the resource $S_{t_P}$, and $S_{T_P}$, the statistical margin according to such approach is:

$$\gamma = \frac{Y - \bar{Y}}{S_T + S_Y}$$

(12)

The given dependence according to its structure is similar to the Cornell safety index, where the normal distribution of the value $\ln T$ is used [18,21]. The difference between the two approaches for determination of the failure probability is in the fact that comparing the pdf $f(y)$ and $f(Y)$ in order to find the value of $\gamma$ by (11) in its denominator the hypotenuse of the rectangular triangle with the catets in the form of SD $S_y$ and $S_Y$ (Fig. 4) is used. When comparing the functions of the technical condition distribution indicators in the denominator (12), the sum of the catets or the SD $S_T$ and $S_Y$ (Fig. 4, right) is substituted. This form is used in the algorithm of the determination of the total strength margin by the partial margin of the influence factors [23]. Here the relationship between formulas (12) and (11) is carried out due to the correction function $a(y/Y)$:
\[ \sqrt{S_y^2 + S_y^2} = (S_y + S_y) \cdot \alpha(y/Y) \]  
(13)

Its value is in the range from 0.707 to 1.0. Providing, that \( \alpha = 0.75 \), then with a stock \( Y/y = 0.25 - 4.0 \) the difference between the results obtained in (12) and (11) will not exceed 10% [24]. The value \( \alpha = 0.707 \) corresponds to the situation of the equality of the SD \( S_y = S_y \) (Fig. 4, right). Then the size of the correction can be set depending on the linear form of the function \( \alpha(y/Y) \):

\[ \alpha = 1 - 0.293 \frac{S_y}{S_y} \]  
(14)

Here the ratio of SD should be more than one. For \( (S_y/S_y) < 1 \), we must take in (14) the inverse relation of the SD \( S_y/S_y \).

Thus, the second approach (the method of the RSI \( \beta_{EX} \)) is more conservative, contributing to decrease the risk and guaranteed security providing. Such interpretation of the failure probability logically follows from the definition of the safety index based on the resource margin (more precisely the lifetime). The described principle of the statistical resource margin determination is fully proved, because the durability SD is sufficiently greater than the operation life SD: \( S_t >> S_t \).

4.2. The Amalgamation of the Resource Safety Indices.

Dimensionless risk index of the system is as follows [15]:

\[ \rho_{\Sigma} = \sum Q_{ik} u_{ik} \]  
(15)

Taking into account that the deterministic damage \( a = t / T_0 = \lambda t \), the exponential distribution law is transformed as \( P_{EXP} = \exp (-a) \). In operational safety strategies, first of all, high-risk elements are considered, for which the deterministic damage, as a rule, does not exceed 0.25. Then, with an error of not more than 5%, you can take \( P_{EXP} = 1 - a \). It follows that \( a = Q \). Combined (9) and (15) get the RSI for systems:

\[ \beta_{EX} = \lg \left( \frac{\sum u_{ik} / 10^{P_{50}}}{} \right)^{-1} \]  
(16)

Guided by the same considerations, we can find the rule for amalgamating the safety indexes for the reliability functions described by the Lindley law \( P_{LND} = (1 + a) P_{EXP} \). Then they are represented in the form \( P_{LND} = (1 + a) (1-a) = 1-a^2 \). From this it follows that \( a^2 = Q \). Then, analogously to (16), we obtain:

\[ \beta_{EPLND} = \lg \left( \frac{\sum u_{ik} / 10^{P_{50}}}{} \right)^{-1} \]  
(17)

5. SEARCH OF THE GENERAL SAFETY INDEX FOR A COMPLEX OF DAMAGING PROCESSES

The first situation of the structural approach of reliability assessment is considered (Fig. 1, a). The effectiveness of the RSI method demonstrated by the example of high-strength M18 bolts for connecting aircraft structural, including chassis wheels (Fig. 7, 8). The detail is a responsible element. For dangerous places bolt where fractures occurs (1, 2, 3, 4, Fig.8) obtained the fatigue models (Table 1). Fractures in places 2 and 3 are typical for imperfect bolts.

The first model is the lifetime general equation (LGE) for lifetime (durability)N in the form:

\[ \lg N = b_0 + m \lg \Delta F + b_{\sigma} R_{\sigma} \]  
(18)

where \( b_0, m, b_{\sigma} \) - coefficients of the model,
\( \Delta F \) is the double amplitude (swing) of the force acting on the bolt in kN,
\( R_{\sigma} \) – Stress ratio.
The second model is the equation for the dispersion of durability (EDD) in the form of:

\[ S_{lgN} = B + k_L (lg N - lg N_A) \],

where \( S_{lgN} \) - SD of the number of cycles to the boundary state, 

\( B, k_L, lgN_A \) – coefficients of the model.

For dangerous places bolt where fractures occurs (1,2,3,4, Fig.8) obtained the parameters fatigue models (Table 1).

**Table 1.** Parameters of fatigue models for dangerous bolt places and their criticality (LGE is obtained for double the amplitude of the forces \( \Delta F \) (kN))

<table>
<thead>
<tr>
<th>dangerous place</th>
<th>LGE</th>
<th>EDD</th>
<th>( u_{ik} )</th>
</tr>
</thead>
<tbody>
<tr>
<td>thread, 1</td>
<td>30.7</td>
<td>-13.2</td>
<td>-2.5</td>
</tr>
<tr>
<td>head surface, 4</td>
<td>14.5</td>
<td>-4.5</td>
<td>-2.5</td>
</tr>
<tr>
<td>fillets, 2,3</td>
<td>21.5</td>
<td>-8.6</td>
<td>-2.5</td>
</tr>
</tbody>
</table>

The task was to find the cyclic lifetime of \( N_P \), which guarantees the safety of PS \( P=0.98 \). For this purpose, according to (9), the value \( \beta_{LGE} = lgN_P \) is determined (Fig. 8). The calculation is performed for a fixed tightening forces \( F \) generated in the body of the bolt tension 0.3 and 0.6 of yield strength. Changing the value of external loading \( \Delta F \) alters stress ratio \( R_{\sigma} \). Therefore, the use of LGE is convenient.

On the basis of the developed algorithm, the diagrams of the initial RSI \( \beta_{LGE} \) were obtained (Fig. 8). Since during construction, the external load is considered as a variative, this diagram is, in essence, a fatigue curve for the random load at the PS = 0.98.
Studies have shown that the factor of variation in external loading of $V_F$ has a more significant effect on the guaranteed lifetime than the tightening force $F$. The diagrams obtained allow 4-10 times longer use of the details than with the forecast by traditional means (Fig.9).

Considering the defined guaranteed lifetime of the object under the system of damaging processes, we can consider the index of criticality $u_k$ as a powerful instrument for the regulation of amalgamated reliability. Such a conclusion follows from the fact that the guaranteed lifetime increases 4-5 times in the transition from the situation $u_k = l$ to the algorithm with the actual calculated $u_k < l$ (Fig.9). That is, in the first situation, the amalgamated RSI is significantly smaller than the mean $\beta_{u_P}$ between the individual indexes: $\beta_{\Sigma P} < \beta_{u_P}$. In the second situation, the principle $\beta_{\Sigma P} \rightarrow \beta_{u_P}$, which can be considered analogous to the above principle $P_z \rightarrow P_{ikmin}$, is formalized.

6. CONCLUSION

The structure of amalgamating formulas, which exclude excessive conservatism when calculating the reliability of the system, is found. New rules for amalgamating based on the risk indicator and on the basis of Lindley's distribution are obtained.

The algorithm for search of the safety index for systems is proved, which ensures the implementation of the principle $\beta_{\Sigma P} \rightarrow \beta_{u_P}$, which can be considered analogous to the principle $P_z \rightarrow P_{ikmin}$. The rule for amalgamating individual indices based on the distribution of Lindley $\beta_{\Sigma P_{LND}}$ is proposed. It is recommended to use it for a large number of (more than 10) critical elements of the system and for multi-site damage. With less number of them it is suggested to use a more usual form of the RSI $\beta_{\Sigma P}$. It well meets the situation of several (4-7) degradation processes on the element of the technical system.

An explanation of the low reliability of the system, which is determined by the rule of multiplication of the PS, is found. Usually it is associated with the factor of mutual influence of elements. In the method of the safety index, the situation $P_z << P_{ikmin}$ is explained by the uncertainty of models. Its influence is offset by the identification of models, among which the criticality of failure $u_k$ is important.

The algorithm for constructing the diagram "RSI of the system $\beta_{\Sigma P}$ - load parameter (in this case $\Delta F$)" is a some alternative to the procedure of summation of the damage. The latter is relevant at the early stages of design, when the uncertainty of the loading to be chosen the spectrum with a wide variation. At the stage of operation, when the loading process is monitored, its variation is substantially reduced. It creates the opportunity not to sum up the damage directly to control the exhaustion of the resource or the remaining lifetime.

REFERENCES


